Homework 4.2 Stat 605, Fall 2018

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Due Thursday, October 4

In class, I mentioned that by making an analogy with inner products of vectors in \mathbb{R}^3 one might interpret $\rho(X, Y)$ as the cosine of the angle between $X - \mathbb{E}X$ and $Y - \mathbb{E}Y$. Therefore, one might say that correlation measures the degree to which the deviations of X and Y from their expectations point in the same direction. The following problem is intended to clarify a subtly different statement about correlation: Correlation measures the strength of the linear relationship between two random variables.

Problem 1 (1+2+2 points). Let $X, Y : (\Omega, \mathscr{F}, \mathbb{P}) \to (\mathbb{R}, \mathscr{B})$ be random variables.

- 1. Write down an L^2 -orthonormal basis for span $\{Y, \mathbf{1}\}$. Here, $\mathbf{1}$ denotes the constant random variable with value 1. Note that span $\{Y, \mathbf{1}\}$ is the set of all linear functions $\alpha Y + \beta$ of Y.
- 2. Write down a formula for the orthogonal projection of X onto span{Y,1}. Try to make the correlation $\rho(X, Y)$ appear in this formula.
- 3. Define

$$d(X, \text{span}\{Y, \mathbf{1}\}) = \min_{Z \in \text{span}\{Y, \mathbf{1}\}} ||X - Z||_2$$

Derive a formula expressing $d(X, \operatorname{span}\{Y, \mathbf{1}\})$ in terms of $\rho(X, Y)$ and $\operatorname{var}(X)$. Note: When statisticians say things about R^2 values and the proportion of the variation in X explained by a linear regression on Y, they are using a version of this formula.

Recall that for $X : (\Omega, \mathscr{F}, \mathbb{P}) \to (U, \mathscr{S})$ a random variable, we define the σ -algebra generated by X to be

$$\sigma(X) = \{X^{-1}(B); B \in \mathscr{S}\}$$

(I believe that this definition first appeared in our discussion of independent random variables.) The σ -algebra $\sigma(X)$ is sometimes called the information contained in X. To see why, work out the following problem.

Problem 2 (2+4 points). For \mathscr{G} a σ -algebra on Ω , we call a map $H : \Omega \to \mathbb{R}$ \mathscr{G} -measurable if $H^{-1}(B) \in \mathscr{G}$ for all $B \in \mathscr{B}$. We say that a function $H : \Omega \to \mathbb{R}$ is \mathscr{G} -simple if it is a \mathscr{G} -measurable simple function.

- 1. Show that the set of all G-measurable maps is the smallest class containing the simple functions and closed under pointwise limits. Note: Most of the work here has been done in class.
- 2. Show that Y is $\sigma(X)$ -measurable if and only if Y = f(X) for some measurable $f : (\mathbb{R}, \mathscr{B}) \to (\mathbb{R}, \mathscr{B})$. In particular, if two random variables contain the same information, then they are functions of each other.

We will see that many properties of a random variable depend only on the information which it contains. For example, $X \perp\!\!\!\perp Y$ implies that $X' \perp\!\!\!\perp Y'$ for any $\sigma(X)$ -measurable X' and $\sigma(Y)$ -measurable Y'. Later, we will see that the conditional expectation $\mathbb{E}[X|Y]$ depends only on X and the information contained in Y.

The results in this problem may possibly be useful below.

Problem 3 (1+1+1 points). A random variable X is said to be a Bernoulli trial with probability of success p if X takes only the values $\{0,1\}$ and $\mathbb{P}[X=1] = p$. Let $\{X_i\}_{i=1}^{\infty}$ be independent Bernoulli trials, and define

$$T = \inf\{m : X_m = 1\}.$$

The distribution of T is called the geometric distribution. Since T takes values in the countable set \mathbb{N} , this distribution is characterized by the probability mass function $f_T : \mathbb{N} \to \mathbb{R}$ by

$$f_T(t) = \mathbb{P}[T=t].$$

- 1. Find a formula expressing $f_T(t)$ in terms of p and t.
- 2. Compute $\mathbb{E}[T]$.
- 3. Compute $\operatorname{var}(T)$.

This next problem is extra credit. If you like, you can use the result below even if you do not turn in a proof. It is helpful to introduce some new terminology here: We say that the functions $f, g : \mathbb{R} \to \mathbb{R}$ are asymptotic in the limit as x tends to c if

$$\lim_{x \to c} \frac{f(x)}{g(x)} = 1.$$

Here, c may be a real number or it may be infinity. When f is asymptotic to g in some limit, we write $f \sim g$. It is important to note that this definition is really only interesting when either f and g both tend to zero in the limit or both tend to infinity in the limit. Otherwise, at least if both limits exist, f is asymptotic to g if and only if the limits are equal.

Problem 4 (1 points). Show that $n \sum_{m=1}^{n} m^{-1} \sim n \log n$ as $n \to \infty$.

Problem 5 (1 point). Let $\{Y_n\}$ be a sequence of random variables and v_n a sequence of real numbers so that $\operatorname{var}(Y_n)/v_n^2 \to 0$ as $n \to \infty$. Show that

$$\frac{Y_n - \mathbb{E}Y_n}{v_n} \xrightarrow{\mathbf{P}} 0.$$

Problem 6 (5 points). Let $\{X_i\}_{i=1}^{\infty}$ be independent random variables, each having the uniform distribution on the finite set $\{1, 2, ..., n\}$. Let

$$\tau_k^n = \inf\{m : \#\{X_1, \dots, X_m\} = k\}$$

be the least index m for which the sequence X_1, \ldots, X_m takes k distinct values. (Here, we use #S to denote the cardinality of the set S.) To make this more concrete, let us suppose that you like watching birds and that you will stop only once you have seen every bird on a list of n birds. If we assume that the birds are observed with equal probability and that the observations are independent, then τ_n^n is a good model for the total number of observations that you will make before stopping. Prove that

$$\frac{\tau_n^n}{n\log n} \stackrel{\mathrm{P}}{\longrightarrow} 1.$$

Hint: Use the results of the previous three problems.